

Abstract

Context: Sparse PCs recovery given a set of data observations of high dimension.

Objective: Study of kernel random matrices of the form $f(\hat{C})$ where \hat{C} is the sample covariance matrix of the data.

Result: f such that $f'(0) = f''(0) = 0$ allows for powerful recovery of sparse PCs.

Model & Assumptions

Data matrix:

$$X \equiv \Sigma_p^{1/2} Z = \left(I_p + \sum_{i=1}^k \omega_i u_i u_i^\top \right)^{1/2} Z \quad (\text{of size } p \times n)$$

- Z random matrix with $\mathcal{N}(0, 1)$ i.i.d. entries.
- $U = [u_1, \dots, u_k] \in \mathbb{R}^{p \times k}$ isometric.

Kernel matrix:

$$F = f\left(\frac{1}{n} X X^\top\right) \stackrel{\text{entry-wise}}{\equiv} \left\{ f\left(\left[\frac{1}{n} X X^\top\right]_{ij}\right) \right\}_{i,j=1}^p$$

(A1) As $n \rightarrow \infty$; $p/n \rightarrow c \in (0, \infty)$ and $\limsup_n \max_i \omega_i < \infty$.

(A2) f is three-times continuously differentiable.

Concentration of Measure

Definition 1. A random vector V is said to be normally ($q = 2$) or exponentially ($q = 1$) concentrated if for any 1-Lipschitz real function $\mathbb{P}\{|f(V) - \mathbb{E}f(V)| \geq t\} \leq C e^{-(t/\sigma)^q} \Rightarrow V \in \mathcal{N}(\sigma)$ or $V \in \mathcal{E}(\sigma)$.

(P1) If $V \in \mathbb{R}^p$ Gaussian then $V \in \mathcal{N}(\sigma)$ [1].

(P2) If $V \in \mathbb{R}$ with finite expectation and $V \in \mathcal{N}(\sigma)$ then $V^2 \in \mathcal{E}(\sigma) + \mathcal{N}(\sigma)$.

General Result: Asymptotic Equivalent

$$\tilde{F} \equiv f(\Sigma_p) + \sum_{k=1}^2 \frac{f^{(k)}(\Sigma_p)}{k} \odot \left[\Sigma_p^{1/2} \left(\frac{1}{n} Z Z^\top - I_p \right) \Sigma_p^{1/2} \right]^{\odot k}$$

Proposition 1. Under (A1) and (A2), for $\eta > 0$

$$F = \tilde{F} + \mathcal{O}_\eta(n^{-\frac{1}{2}})$$

where $X = \mathcal{O}_\eta^m(n^{-\alpha})$ means $\mathbb{P}\{\|X\|_{op} \geq C n^{-\alpha} \eta^{-\frac{1}{2m}}\} \leq \eta$.

Sketch of Proof:

- Taylor expansion of F around $f(\Sigma_p)$.
- Control of the operator norm of $f^{(3)}(\xi^n) \odot \left[\Sigma_p^{1/2} \left(\frac{1}{n} Z Z^\top - I_p \right) \Sigma_p^{1/2} \right]^{\odot 3}$.
- Key lemma: $g(Z) = v^\top (Z Z^\top / n - I_p) v \in \mathcal{E}(n^{-\frac{1}{2}}) + \mathcal{N}(n^{-\frac{1}{2}})$

$$v^\top Z \xrightarrow{(P1)} \|v^\top Z / \sqrt{n}\| \in \mathcal{N}(n^{-\frac{1}{2}}) \xrightarrow{(P2)} \|v^\top Z / \sqrt{n}\|^2 \in \mathcal{E}(n^{-\frac{1}{2}}) + \mathcal{N}(n^{-\frac{1}{2}}).$$

Sparse Matrices

Notations:

- $\mathcal{A}(M) = \{1_{M_{ij} \neq 0}\}_{i,j=1}^p$ defines a graph \mathcal{G}_p .
- $\mathcal{C}_p(k)$ the set of closed walks of length k on \mathcal{G}_p .

Definition 2 (ε -sparse matrices [2]). M is ε -sparse if \mathcal{G}_p satisfies, for all $k \in 2\mathbb{N}$

$$|\mathcal{C}_p(k)| \leq C_k p^{\varepsilon(k-1)+1}$$

Example:

$$M = \begin{pmatrix} 1 & \frac{1}{\sqrt{p}} & \frac{1}{\sqrt{p}} & \frac{1}{\sqrt{p}} & \frac{1}{\sqrt{p}} \\ \frac{1}{\sqrt{p}} & 1 & 0 & 0 & 0 \\ \frac{1}{\sqrt{p}} & 0 & 1 & 0 & 0 \\ \frac{1}{\sqrt{p}} & 0 & 0 & 1 & 0 \\ \frac{1}{\sqrt{p}} & 0 & 0 & 0 & 1 \end{pmatrix} \xrightarrow{\text{Graph representation}} \begin{array}{c} \circ \\ / \quad \backslash \\ \circ \quad \circ \\ / \quad \backslash \\ \circ \quad \circ \end{array}$$

(P3) If M is ε -sparse and $f(0) = 0$, then $f(M)$ is also ε -sparse.

(P4) If M is ε -sparse with bounded entries then $\forall k \in 2\mathbb{N} \quad \|M\|_{op} = \mathcal{O}(p^{\varepsilon(1-1/k)+1/k})$.

Special Case: Sparse PCA

(A3) Σ_p is $\frac{1}{2+\mu}$ -sparse for $\mu > 0$.

Proposition 2. Under (A1), (A2) and (A3), for all $\varepsilon \in (0, \frac{\mu}{2(3+2\mu)})$

$$F = f(\Sigma_p) + \mathcal{O}_\eta^{1/\varepsilon} \left(n^{\frac{\mu}{2(2+\mu)} + \varepsilon(2 - \frac{1}{2+\mu})} \right) \text{ s.t. } f'(0) = f''(0) = 0$$

Sketch of Proof:

Control the operator norm of $f^{(k)}(\Sigma_p) \odot \left[\Sigma_p^{1/2} \left(\frac{1}{n} Z Z^\top - I_p \right) \Sigma_p^{1/2} \right]^{\odot k}$

- $f^{(k)}(\Sigma_p)$ is ε -sparse by (P3), its operator norm is controlled by (P4).
- $\left[\Sigma_p^{1/2} \left(\frac{1}{n} Z Z^\top - I_p \right) \Sigma_p^{1/2} \right]^{\odot k}$ has entries of order $\mathcal{O}(n^{-k/2})$.

Example of function with $f'(0) = f''(0) = 0$

Remark 1. For the spiked model

$$\Sigma_p = I_p + \sum_{i=1}^k \omega_i u_i u_i^\top \Rightarrow \Sigma_p \text{ 0-sparse}$$

$\forall i, u_i$ sparse $\Rightarrow F = f(\Sigma_p) + \mathcal{O}_\eta^{1/\varepsilon} (n^{-\frac{1}{2}+2\varepsilon})$

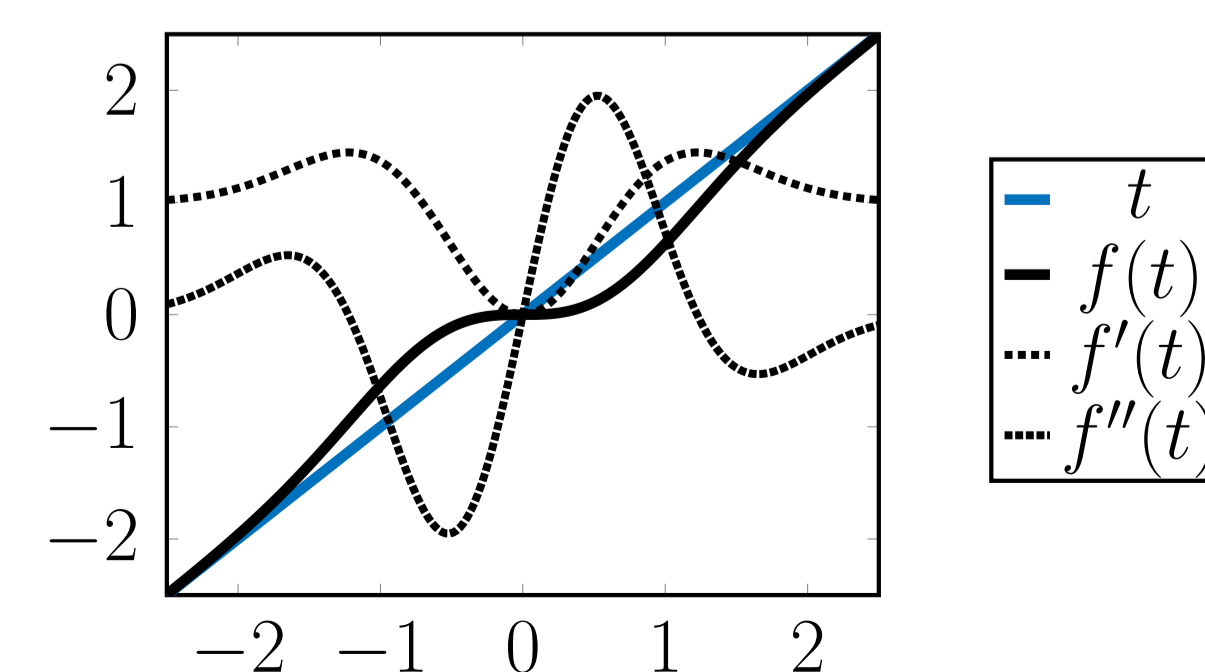
\Rightarrow Choose f such that $f'(0) = f''(0) = 0$.

$\Rightarrow f(t) \approx t$ for large values of t .

For $\alpha > 0$

$$f(t) \equiv t(1 - e^{-\alpha t^2})$$

$$f'(t) = 1 + e^{-\alpha t^2} (2\alpha t^2 - 1) \Rightarrow f'(0) = 0, \\ f''(t) = -2\alpha t e^{-\alpha t^2} (2\alpha t^2 - 3) \Rightarrow f''(0) = 0.$$



Experiments

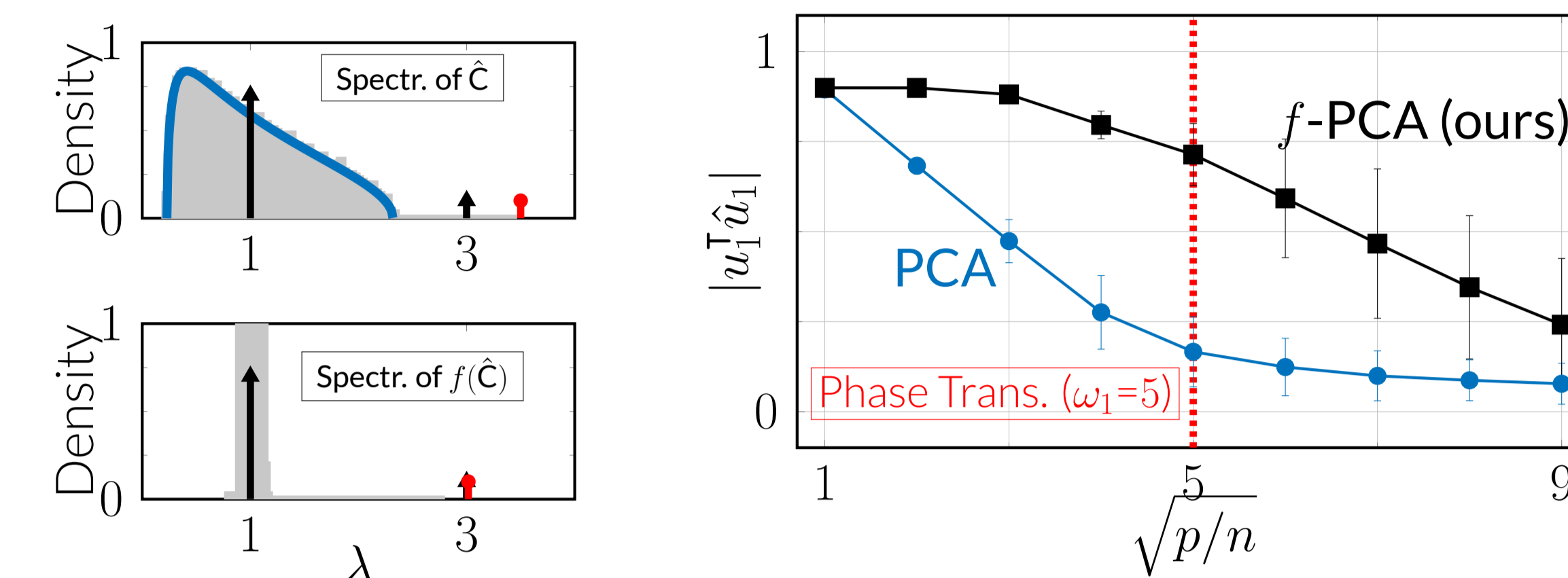


Figure 1. (Left) Spectrum of \hat{C} (up) and $f(\hat{C})$ (bottom) for $p = 2048$ and $n = 7500$. Limiting Marcenko-Pastur density versus spectrum of Σ_p . (Right) Alignment between estimated PC and GT.

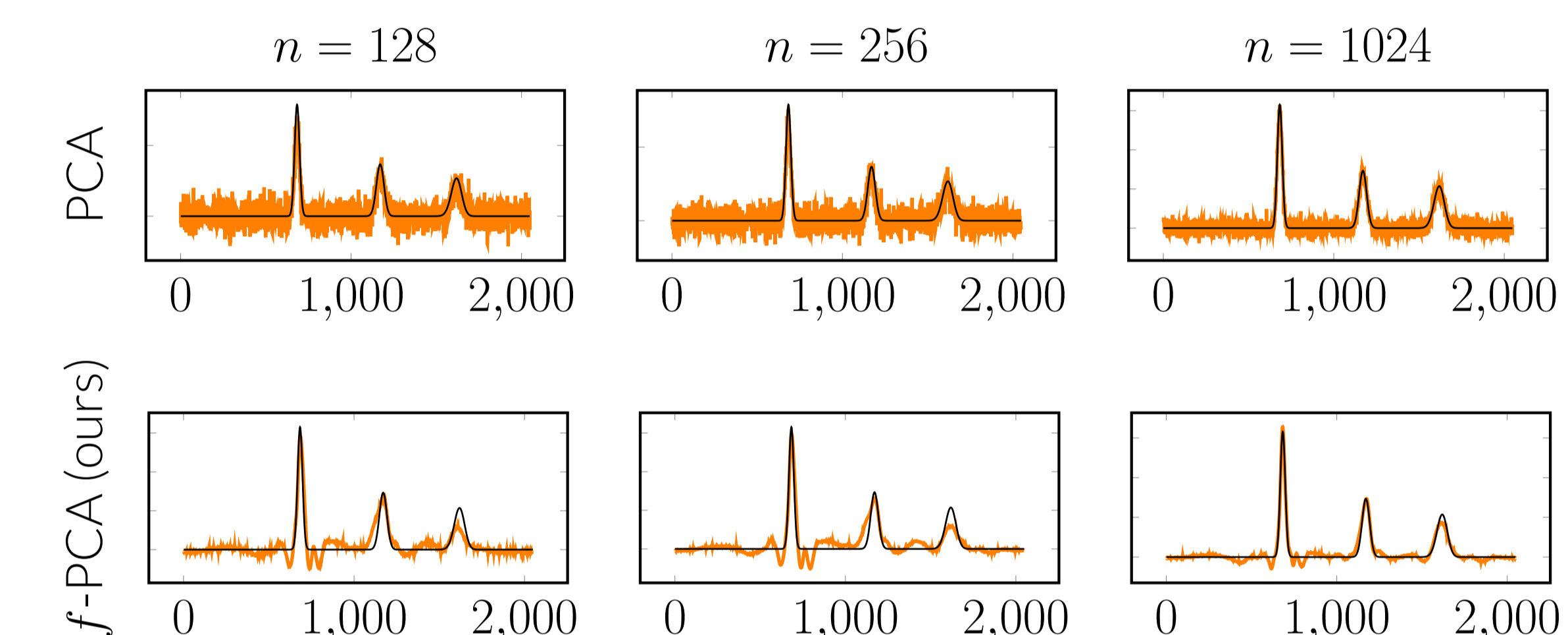


Figure 2. PC recovery by PCA (up) and our method (down). $p = 2048$ and $\omega_1 = 5$.

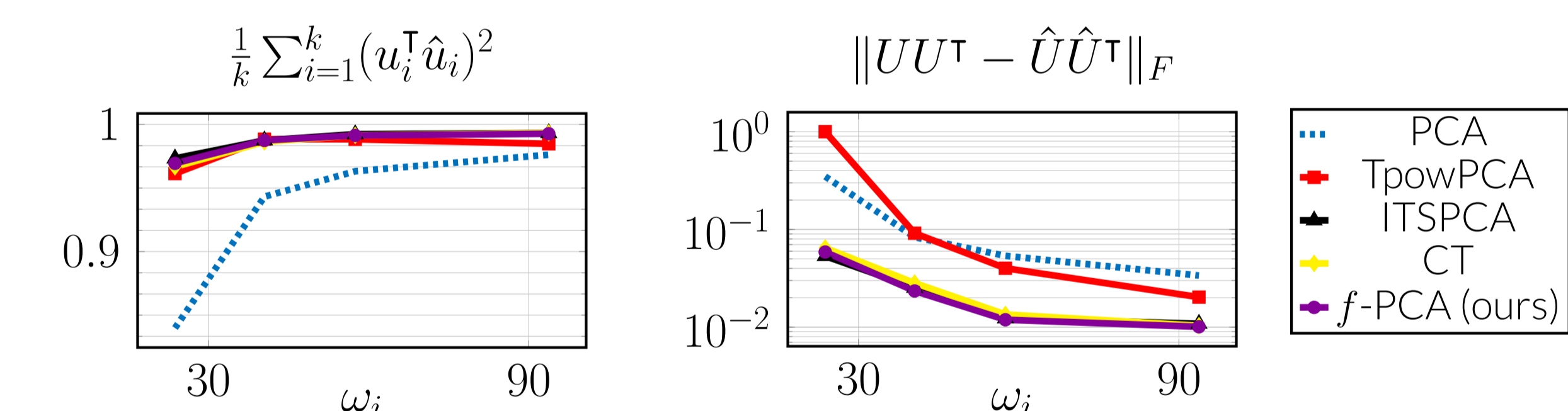


Figure 3. Performances of standard PCA, different state-of-the-art sparse PCA methods and our method. $p = 2048$, $n = 1024$ and $k = 4$.

Perspectives

- Consider the non-trivial setting of Cheng *et al.* [3].
- Estimate optimal hyper-parameter choices based on this setting.

References

- Terence Tao. Topics in random matrix theory, volume 132. American Mathematical Society Providence, RI, 2012.
- Noureddine El Karoui. Operator norm consistent estimation of large-dimensional sparse covariance matrices. The Annals of Statistics, 2008.
- Xiuyuan Cheng and Amit Singer. The spectrum of random inner-product kernel matrices. Random Matrices: Theory and Applications, 2013.